# UNIVERSITY OF SURREY<sup>©</sup>

B. Sc. Honours Courses in Mathematical Studies

## Level HE3 Examination

Module MS331 BAYESIAN STATISTICS

Time allowed - 2 hours

Autumn Semester 2004

Attempt THREE questions. If any candidate attempts more than THREE questions only the best THREE solutions will be taken into account. A formula sheet and Cambridge Statistical Tables will be provided.

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(a) A random sample  $t_1, \ldots t_n$  of lifetimes of batteries follows an exponential distribution with unknown mean  $\theta^{-1} > 0$ . Obtain the likelihood function of the sample and show that the gamma (a,b) distributions with densities

$$p(\theta) = \frac{b^a \theta^{a-1} e^{-b\theta}}{\Gamma(a)}, \ a, b > 0,$$

form a conjugate family. Obtain the posterior distribution of  $\theta$  when the prior distribution is a member of this family. Identify this distribution. [6]

(b) Obtain the predictive distribution of the lifetime T of another battery and show that

$$Pr(T \ge t | t_1, \dots, t_n) = \left(\frac{b+s}{b+s+t}\right)^{a+n}$$

where  $s = \sum_{i} t_{i}$ . The following lifetimes in months of batteries were recorded:

Find the probability that the lifetime of another battery exceeds 8 months, given the prior distribution of  $\theta$  is gamma (2,10). [13]

(c) It is required to obtain the predictive density of the total lifetime Y of m future batteries. What is the distribution of Y given θ? Obtain the predictive density of Y up to a constant of proportionality.

(a) The results of k experiments on female rats yield the numbers  $y_i$  of rats who developed an endrometrial stromal polyp out of the total number  $n_i$ of rats in experiment i for i = 1, ..., k. The data are modelled by assuming that the  $y_i$  are independent binomial random variables with parameters  $n_i, \theta_i$ , respectively, where  $\theta_i$  is the probability of a rat developing a polyp in the *i*th experiment. Write  $\underline{y} = (y_1, y_2, ..., y_k), \underline{\theta} = (\theta_1, ..., \theta_k)$ . Show that the likelihood function of the data is given by

$$p(\underline{y}|\underline{\theta}) \propto \prod_{i=1}^{k} \theta_i^{y_i} (1-\theta_i)^{n_i-y_i}$$

(b) Now assume that the  $\theta_i$  are an independent sample from a beta  $(\alpha, \beta)$  distribution. That is,

$$p(\underline{\theta}|\alpha,\beta) = \prod_{i=1}^{k} \left\{ \frac{\theta_i^{\alpha-1} (1-\theta_i)^{\beta-1}}{B(\alpha,\beta)} \right\}$$

where B(.,.) is the beta function. Further assume a vague prior for  $(\alpha,\beta)$  of the form  $p(\alpha,\beta) \propto (\alpha\beta)^{-1}$ . Show that the full posterior density  $p(\underline{\theta},\alpha,\beta|\underline{y})$  is proportional to  $(\alpha\beta)^{-1}p(\underline{\theta}|\alpha,\beta)p(\underline{y}|\theta)$ . Hence write down an expression for  $p(\underline{\theta},\alpha,\beta|\underline{y})$ . (Do not attempt to calculate the constant of proportionality.)

(c) Show that the joint marginal posterior density of  $(\alpha, \beta)$  is given by

$$p(\alpha,\beta|\underline{y}) \propto \frac{\prod_{i=1}^{k} B(\alpha+y_i,\beta+n_i-y_i)}{\alpha\beta\{B(\alpha,\beta)\}^k}$$

Suppose that the parameter  $\mu = \frac{\alpha}{\alpha + \beta}$  is of primary interest. Deduce the form of the joint posterior density of  $(\mu, \phi)$ , where  $\phi = \alpha + \beta$ . Suggest a method for evaluating the marginal posterior distribution of  $\mu$ . [16]

#### Question 3

(a) Define a  $100(1 - \alpha)\%$  credible interval and a  $100(1 - \alpha)\%$  Highest Posterior Density (HPD) interval for an unknown parameter  $\theta$ . Prove that a  $100(1 - \alpha)\%$  HPD interval is the shortest possible credible interval, assuming that the posterior distribution has a unimodal density.

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[3]

[6]

[8]

(b) Observations are assumed to come from a Cauchy distribution with pdf

$$p(x|\theta) = \pi^{-1} [1 + (x - \theta)^2]^{-1} \quad -\infty < x < \infty$$

The prior density for  $\theta$  is assumed to be proportional to a constant. Thus the posterior density for  $\theta$  is given by

$$p(\theta|\underline{x}) \propto \prod_{i=1}^{n} [1 + (x_i - \theta)^2]^{-1}$$

Suppose there are 5 observations 2.1, 4.9, 3.5, 2.8 and 4.1. Find the missing lines in the following table corresponding to  $\theta = 4.0, 4.5, 5.0$  where

$$H(\theta) = 10^{5} [1 + (2.1 - \theta)^{2}]^{-1} [1 + (4.9 - \theta)^{2}]^{-1} \dots [1 + (4.1 - \theta)^{2}]^{-1}$$

and the integrals are calculated using Simpson's rule.

| $\theta$ | $H(\theta)$ | $\int_{-\infty}^{\theta} H(t) dt$ | $p(\theta \underline{x})$ | $\int_{-\infty}^{\theta} p(t \underline{x}) dt$ |
|----------|-------------|-----------------------------------|---------------------------|---|
| 0.0      | 0           | 0                                 | 0                         | 0   |
| 0.5      | 2           |                                   | .000                      |   |
| 1.0      | 9           | 2.8                               | .000                      | .000  |
| 1.5      | 56          |                                   | .006                      |   |
| 2.0      | 365         | 102.5                             | .042                      | .012  |
| 2.5      | 1643        |                                   | .187                      |   |
| 3.0      | 4171        | 1953.8                            | .475                      | .223  |
| 3.5      | 5632        |                                   | .642                      |   |
| 4.0      |             |                                   |                           |   |
| 4.5      |             |                                   |                           |   |
| 5.0      |             |                                   |                           |   |
| 5.5      | 48          |                                   | .005                      |   |
| 6.0      | 7           | 8777.5                            | .000                      | 1.000   |
| 6.5      | 1           |                                   | .000                      |   |
| 7.0      | 0           | 8779.3                            | .000                      | 1.000   |

Explain how you would amend the table if the prior for  $\theta$  was normal with mean  $\mu$  and variance  $\sigma^2$ .

Why is numerical integration important for the Bayesian approach to statistics?

[3]

[4]

[10]

(a) A three-stage linear model is given by

where  $\underline{y}$  is an  $n \times 1$  vector,  $\underline{\theta}_1$  is a  $p_1 \times 1$  vector,  $\underline{\theta}_2$  is a  $p_2 \times 1$  vector,  $A_1$ ,  $A_2, C_1, C_2, C_3$  and  $\mu$  are known.

Using the results for the two stage model given below show that the posterior distribution of  $\underline{\theta}_1$  is  $N(D\underline{d}, D)$  where

$$D^{-1} = A_1^T C_1^{-1} A_1 + \{C_2 + A_2 C_3 A_2^T\}^{-1},$$
  

$$\underline{d} = A_1^T C_1^{-1} \underline{y} + \{C_2 + A_2 C_3 A_2^T\}^{-1} A_2 \underline{\mu}.$$

Use the matrix lemma given below to find the corresponding result when  $C_3^{-1} \rightarrow 0$ .

[Note: The two-stage linear model is given by

$$\underline{y}|\underline{\theta}_1, C_1 \sim N(A_1\underline{\theta}_1, C_1)$$
  
$$\underline{\theta}_1|\mu, C_2 \sim N(\mu, C_2),$$

where  $A_1$ ,  $C_1$ ,  $C_2$  and  $\underline{\mu}$  are known. The marginal distribution of  $\underline{y}$  is  $N(A_1\underline{\mu}, C_1 + A_1C_2A_1^T)$  and the posterior distribution of  $\theta_1$  is N(Bb, B) where

$$B^{-1} = A_1^T C_1^{-1} A_1 + C_2^{-1}$$
  

$$\underline{b} = A_1^T C_1^{-1} \underline{y} + C_2^{-1} \underline{\mu}$$

Matrix lemma: for any matrices  $A_1$ ,  $C_1$ ,  $C_2$  of appropriate dimensions for which the inverses stated in the result exist we have

$$\{C_1 + A_1 C_2 A_1^T\}^{-1} = C_1^{-1} - C_1^{-1} A_1 (A_1^T C_1^{-1} A_1 + C_2^{-1})^{-1} A_1^T C_1^{-1}\}.$$

(b) Consider a simple linear regression model  $y_i | \alpha, \beta \sim N(\alpha + \beta(x_i - \bar{x}), \sigma^2)$ for i = 1, 2, ..., n where  $\alpha \sim N(0, \sigma^2), \beta \sim N(2, \sigma^2)$  and  $Cov(\alpha, \beta) = 0$ . Write this as a two stage normal linear model. If the values of  $y_i$  and  $x_i$ are as given below, find the posterior distributions of  $\alpha$  and  $\beta$  in terms of  $\sigma$ . [16]

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[9]

(a) It is desired to estimate the posterior means of three parameters,  $\theta$ ,  $\delta$  and  $\phi$ . The full conditional distributions,

$$p(\theta|\delta,\phi,\text{data}), \quad p(\delta|\phi,\theta,\text{data}), \quad p(\phi|\theta,\delta,\text{data})$$

are known. Explain how the posterior means may be estimated using the Gibbs sampling method. [10]

(b) The brightness (J) of a comet may be modelled by a Pareto distribution with density

$$p(J \mid \alpha) = \frac{\alpha J_x^{\alpha}}{J^{\alpha+1}} \qquad J > J_x$$

where  $J_x$  is the lower limit of brightness for the sample considered and  $\alpha$  is a brightness index.

(i) If a sample of brightness values  $J_1, J_2, \ldots, J_n$  for *n* comets is available show that the maximum likelihood estimate of  $\alpha$  is given by

$$\hat{\alpha} = \frac{n}{\sum_{i=1}^{n} \log_e(J_i/J_x)}.$$

[4]

[4]

- (ii) If the prior distribution for  $\alpha$  is taken as Gamma(a, b) find the posterior distribution in terms of a, b, n and  $\hat{\alpha}$ .
- (iii) Three groups of comets of sizes  $n_1, n_2$  and  $n_3$  have brightness parameters  $\alpha_1, \alpha_2$  and  $\alpha_3$ . The prior distribution for each  $\alpha_i$  is taken as Gamma(a, b). For physical reasons the brightness parameters must be ordered  $\alpha_1 < \alpha_2 < \alpha_3$ . Explain how the posterior means of the parameters can be estimated by adapting the Gibbs sampling algorithm you have described in part (a). [7]

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