

1. Suppose  $Y_1, Y_2, Y_3, \dots$  are independent Bernoulli random variables with respective success probabilities  $\frac{1}{2}, \frac{1}{4}, \frac{1}{8}, \dots$ . That is, for  $i = 1, 2, \dots$  the probability mass function of  $Y_i$  is given by

$$P(Y_i = 0) = 1 - \left(\frac{1}{2}\right)^i,$$

$$P(Y_i = 1) = \left(\frac{1}{2}\right)^i.$$

For  $n = 1, 2, \dots$  define  $S_n = Y_1 + \dots + Y_n$ .

- (a) For  $n = 3$ , compute the probability mass function of  $S_3$ . [6 marks]
- (b) Find an expression for the probability  $P(S_n = n)$  for  $n = 1, 2, \dots$  [2 marks]
- (c) For large  $n$ , can the Central Limit Theorem be used to approximate the distribution of  $S_n$ ? Explain your answer. [2 marks]
- (d) Find  $E[S_n]$  and  $\text{Var}[S_n]$ .

[ You may use without proof the result that

$$\sum_{i=0}^n x^i = \frac{1 - x^{n+1}}{1 - x}. \quad ]$$

[6 marks]

- (e) Defining the sample mean  $\bar{Y}_n = S_n/n$ , show that  $\lim_{n \rightarrow \infty} E[\bar{Y}_n] = 0$  and find the value of  $\lim_{n \rightarrow \infty} \text{Var}[\bar{Y}_n]$ . Comment on why these limiting values are intuitively reasonable. [4 marks]

2. (a) Suppose the random variable  $T$  is exponentially distributed with parameter  $\lambda$ , so that  $T$  has distribution function

$$F_T(t) = \begin{cases} 0 & t < 0, \\ 1 - e^{-\lambda t} & t \geq 0. \end{cases}$$

Define the *memoryless property*, explain the intuitive meaning of this property, and show that the distribution of  $T$  has the memoryless property. [4 marks]

- (b) The random variable  $X$  is said to follow the Weibull distribution with parameters  $\beta > 0$ ,  $\theta > 0$  if the distribution function of  $X$  is given by

$$F_X(x) = \begin{cases} 0 & x < 0, \\ 1 - \exp\left\{-\left(\frac{x}{\theta}\right)^\beta\right\} & x \geq 0. \end{cases}$$

Derive an expression for the probability density function  $f_X(x)$  of  $X$ . [2 marks]

Suppose the random variable  $X$  follows the Weibull distribution with parameters  $\beta, \theta$ . For  $a, b > 0$  find an expression for  $P(X > a + b \mid X > a)$ . [3 marks]

Show that in the case  $\beta = 2$ , the distribution of  $X$  does not possess the memoryless property. [3 marks]

For which values of the parameters  $\beta, \theta$ , if any, does the Weibull distribution possess the memoryless property? [1 mark]

- (c) Suppose that the lifetime  $Y$  (in years) of a particular electrical component follows the Weibull distribution with parameters  $\beta = 2$ ,  $\theta = 1$ . Treating  $b$  as a fixed constant, sketch the graphs of the functions  $g(a)$  and  $h(a)$  defined by

$$g(a) = P(Y > a + b \mid Y > a),$$

$$h(a) = \frac{P(Y > a + b \mid Y > a)}{P(Y > b)}.$$

Give your interpretation of the shape of the graphs of  $g(a)$  and  $h(a)$ . [7 marks]

3. (a) Suppose that  $X$  is a continuous random variable with probability density function  $f_X(x)$ , and that the random variable  $Y$  is defined by  $Y = g(X)$  for some strictly monotonic function  $g(x)$ . Show that the probability density function of  $Y$  is given by

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right|.$$

[5 marks]

- (b) Suppose  $X$  is a continuous random variable with density

$$f(x) = \begin{cases} x^3/4 & 0 < x < 2, \\ 0 & \text{otherwise.} \end{cases}$$

- (i) Find the mean, median and mode of the distribution of  $X$ . [3 marks]  
(ii) Find the cumulative distribution function  $F(x)$  of  $X$ . [2 marks]  
(iii) Find the cumulative distribution function and the probability density function of the random variable  $Y$  defined by  $Y = \sqrt{(X + 2)}/2$ . What is the range of  $Y$ ?

[7 marks]

Find the value of  $E[Y]$ .

[3 marks]

4. (a) Give formulae defining the *covariance* and the *correlation* of two random variables  $X$  and  $Y$ . Explain how correlation values may be interpreted. [4 marks]
- (b) Suppose that  $X, Y$  are continuous random variables with joint density function

$$f_{X,Y}(x, y) = \frac{e^{-y}}{2y}, \quad y > 0, \quad -y < x < y.$$

- (i) Draw the region of non-zero density. [2 marks]
- (ii) Find the marginal density  $f_Y(y)$  of  $Y$  and for  $y > 0$  find the conditional density  $f_{X|Y}(x | y)$  of  $X$  given  $Y = y$ . [5 marks]  
For each of these two distributions, give the standard name of the distribution, specifying any parameter values. [2 marks]
- (iii) Find  $E[X]$  and  $E[Y]$ . [3 marks]
- (iv) Find  $E[XY]$  and hence find the covariance  $\text{Cov}[X, Y]$ . [3 marks]
- (v) Are  $X$  and  $Y$  independent? Explain your answer. [1 mark]

5. Suppose that  $X, Y$  are continuous random variables with joint density function

$$f_{X,Y}(x, y) = \begin{cases} (2/\pi) \exp \{-(x^2 + y^2)/2\} & \text{for } x, y > 0, \\ 0 & \text{otherwise.} \end{cases}$$

Define the random variables  $U, V$  by

$$U = 2X + Y, \quad V = 3X/Y.$$

- (a) Find the joint density  $f_{U,V}(u, v)$ . Sketch the region where the random vector  $(U, V)$  has positive density. [13 marks]
- (b) Find the marginal density of  $V$ . [7 marks]

[ For part (b), you may use without proof the result that for  $A > 0$ ,

$$\int_0^\infty ue^{-Au^2} du = \frac{1}{2A}. \quad ]$$

6. (a) For any random variable  $X$  the moment generating function  $M_X(t)$  is defined by  $M_X(t) = E[e^{tX}]$ .

Show that (i)  $E[X] = M'_X(0)$  and (ii)  $\text{Var}[X] = M''_X(0) - (M'_X(0))^2$ , where  $M'_X(t)$  and  $M''_X(t)$  denote the first and second derivatives, respectively, of  $M_X(t)$  with respect to  $t$ . [2 marks]

Write down an expression for  $E[X^n]$ , where  $n = 1, 2, \dots$ , in terms of derivatives of the moment generating function  $M_X(t)$ . [1 mark]

Defining the random variable  $Y$  to be  $Y = a + bX$ , where  $a, b$  are non-random, derive an expression for the moment generating function  $M_Y(t)$  in terms of  $M_X$ . [2 marks]

- (b) For any random variable  $X$ , the cumulant generating function  $K_X(t)$  is defined by  $K_X(t) = \ln(M_X(t))$ , where  $M_X(t)$  is the moment generating function defined in part (a) above. Define the cumulants  $\kappa_1, \kappa_2, \dots$  of the distribution.

[1 mark]

What can be said about the relationship between cumulants and moments of a distribution? [2 marks]

- (c) Suppose  $Z$  follows the standard normal distribution,  $Z \sim N(0, 1)$ , so that  $Z$  has probability density function

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \quad -\infty < z < \infty.$$

Show that the moment generating function of  $Z$  is given by  $M_Z(t) = e^{t^2/2}$ . [7 marks]

Hence derive an expression for the moment generating function  $M_X(t)$  of the normal random variable  $X$  with mean  $\mu$  and variance  $\sigma^2$ , and a corresponding expression for the cumulant generating function  $K_X(t)$  defined by  $K_X(t) = \ln(M_X(t))$ . [3 marks]

Based on your expression for  $K_X(t)$ , what can be said about the first, second, and higher cumulants of the  $N(\mu, \sigma^2)$  distribution? To what extent do these statements generalise to distributions other than the normal distribution?

[2 marks]

7. (a) Suppose the random variable  $X$  has the  $\chi_n^2$  distribution for some positive integer  $n$ , so that the probability density function of  $X$  is given by

$$f_X(x) = \frac{1}{2^{n/2} \Gamma\left(\frac{n}{2}\right)} x^{\frac{n}{2}-1} e^{-x/2} \quad \text{for } x \geq 0.$$

Show that the random variable  $U = \sqrt{X/n}$  has probability density function

$$f_U(u) = \frac{n^{n/2}}{2^{(n/2)-1} \Gamma\left(\frac{n}{2}\right)} u^{n-1} e^{-nu^2/2} \quad \text{for } u \geq 0.$$

[7 marks]

- (b) Suppose that  $Y$  is a standard normal random variable, so that  $Y$  has probability density function

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} e^{-y^2/2} \quad \text{for } -\infty < y < \infty.$$

Suppose further that  $Y$  is independent of the variable  $X$  of part (a) above, and define the random variable  $T$  by

$$T = \frac{Y}{\sqrt{X/n}}.$$

Derive an expression for the probability density function  $f_T(t)$  of  $T$ .

[ You may use without proof the results that

- (i) the quotient  $T = Y/U$  of independent continuous random variables  $Y, U$  has density function

$$f_T(t) = \int_{-\infty}^{\infty} |u| f_U(u) f_Y(ut) du;$$

- (ii) for any  $\alpha, \lambda > 0$ ,

$$\int_0^{\infty} x^{\alpha-1} e^{-\lambda x} dx = \frac{\Gamma(\alpha)}{\lambda^\alpha}. \quad ]$$

[13 marks]

