

Bayesian classification

- 1 Discriminative classification
- 2 Bayesian logistic regression & Laplace approximation
- 3 Generative classification

Likelihood model

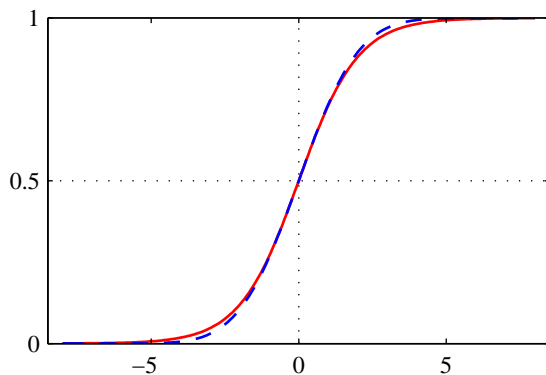
- **Two class** (binary) classification, **discriminative** approach: need model for $p(\mathcal{C}_1|\mathbf{x}, \mathbf{w}) = 1 - p(\mathcal{C}_2|\mathbf{x}, \mathbf{w})$
- Keep this 'almost' linear in parameter vector \mathbf{w} :

$$p(\mathcal{C}_1|\mathbf{x}, \mathbf{w}) = \sigma(\mathbf{w}^T \phi(\mathbf{x})), \quad \sigma(a) = 1/(1 + e^{-a})$$

- $\sigma(a)$ = logistic sigmoid, 'squashing' or '**activation**' function
- Inverse: logit $a = \ln(\sigma/(1 - \sigma))$, 'link' function
- Model known as 'logistic regression' (but it's classification!)
- Other choices for $\sigma(a)$ are possible, e.g. inverse probit

$$\sigma(a) = \int_{-\infty}^a \mathcal{N}(\theta|0, 1)d\theta$$

Activation functions



Red: logistic sigmoid; blue: inverse probit

Likelihood model

- Represent class \mathcal{C}_1 as $t = 1$, \mathcal{C}_2 as $t = 0$, then

$$p(\mathbf{t}|\mathbf{w}) = \prod_{n=1}^N y_n^{t_n} (1-y_n)^{1-t_n}, \quad y_n = p(\mathcal{C}_1|\mathbf{x}_n, \mathbf{w}) = \sigma(\mathbf{w}^T \phi(\mathbf{x}_n))$$

- Maximum likelihood minimizes **cross-entropy** error function

$$E(\mathbf{w}) = -\ln p(\mathbf{t}|\mathbf{w}) = -\sum_n [t_n \ln y_n + (1-t_n) \ln(1-y_n)]$$

- Gradient:

$$\nabla E(\mathbf{w}) = \sum_n (y_n - t_n) \phi(\mathbf{x}_n) = \mathbf{\Phi}^T (\mathbf{y} - \mathbf{t})$$

Likelihood model (2)

- Hessian of E :

$$\nabla\nabla E(\mathbf{w}) = \sum_n y_n(1 - y_n)\phi(\mathbf{x}_n)\phi(\mathbf{x}_n)^T = \mathbf{\Phi}^T \mathbf{R} \mathbf{\Phi}$$

with \mathbf{R} = diagonal matrix, $R_{nn} = y_n(1 - y_n)$

- Positive definite $\Rightarrow E$ is convex, only has a single minimum
- So $p(\mathbf{t}|\mathbf{w})$ is **log-concave**, only has a **single maximum**
- Can be found efficiently numerically (iterative reweighted least squares)

Generalizations

- Allowing **labelling noise**:

$$\begin{aligned} p(\mathcal{C}_1|\mathbf{x}, \mathbf{w}) &= (1 - \epsilon)\sigma(\mathbf{w}^T \phi(\mathbf{x})) + \epsilon[1 - \sigma(\mathbf{w}^T \phi(\mathbf{x}))] \\ &= \epsilon + (1 - 2\epsilon)\sigma(\mathbf{w}^T \phi(\mathbf{x})) \end{aligned}$$

- Classification into **$K > 2$ classes**: use 'softmax'

$$p(\mathcal{C}_k|\mathbf{x}, \mathbf{w}_1 \dots \mathbf{w}_K) = \frac{\exp(a_k)}{\sum_j \exp(a_j)}, \quad a_k = \mathbf{w}_k^T \phi(\mathbf{x})$$

- Likelihood for 1-of- K coding \mathbf{t}_n is then

$$p(\mathbf{t}_1 \dots \mathbf{t}_N | \mathbf{w}_1 \dots \mathbf{w}_K) = \prod_{n=1}^N \prod_{k=1}^K y_{nk}^{t_{nk}}$$

with $y_{nk} = \exp(a_{nk}) / \sum_j \exp(a_{nj})$ and $a_{nk} = \mathbf{w}_k^T \phi(\mathbf{x}_n)$

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Prior and posterior

- Need to put a prior on \mathbf{w} ; could choose as for linear regression
 $p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{0}, \alpha^{-1}\mathbf{I})$
- Gives for posterior $p(\mathbf{w}|\mathbf{t}) \propto p(\mathbf{t}|\mathbf{w})p(\mathbf{w})$

$$\ln p(\mathbf{w}|\mathbf{t}) = -E(\mathbf{w}) + \text{const.}$$

$$E(\mathbf{w}) = \frac{\alpha}{2} \mathbf{w}^T \mathbf{w} - \sum_n [t_n \ln y_n + (1 - t_n) \ln(1 - y_n)]$$

- Need to normalize and then integrate to get predictions

$$p(\mathcal{C}_1|\mathbf{x}, \mathbf{t}) = \int p(\mathcal{C}_1|\mathbf{x}, \mathbf{w})p(\mathbf{w}|\mathbf{t})d\mathbf{w}$$

- Not a Gaussian integral – but $p(\mathbf{w}|\mathbf{t})$ has a single maximum
- So **approximate** by a Gaussian around this maximum:
Laplace approximation

Laplace approximation

In one dimension

- Consider a generic $p(w) = \exp[-E(w)]/Z$, $Z =$ normalization constant ('partition function')
- If $p(w)$ has a single maximum at w_0 , can expand around there:

$$E(w) \approx E(w_0) + \frac{1}{2}E''(w_0)(w - w_0)^2$$

- Gives Gaussian approximation for $p(w)$:

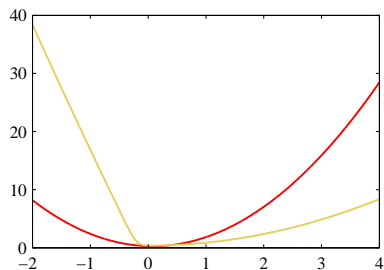
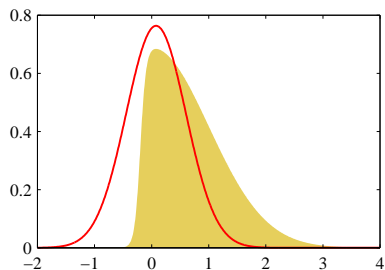
$$p(w) \approx q(w) = \frac{e^{-E(w_0)}}{Z} e^{-\frac{E''(w_0)}{2}(w-w_0)^2} = \mathcal{N}(w|w_0, 1/E''(w_0))$$

- Approximation for Z :

$$Z = e^{-E(w_0)} (2\pi)^{1/2} [E''(w_0)]^{-1/2}$$

Laplace approximation

Illustration



Laplace approximation

In M dimensions

- Consider again $p(\mathbf{w}) = \exp[-E(\mathbf{w})]/Z$
- If $p(\mathbf{w})$ has a single maximum at \mathbf{w}_0 , can expand around there:

$$E(\mathbf{w}) \approx E(\mathbf{w}_0) + \frac{1}{2}(\mathbf{w} - \mathbf{w}_0)^T \mathbf{A}(\mathbf{w} - \mathbf{w}_0)$$

where $\mathbf{A} = \nabla \nabla E(\mathbf{w})|_{\mathbf{w}=\mathbf{w}_0} = \text{Hessian at minimum of } E$

- Gives **Gaussian approximation** for $p(\mathbf{w})$:

$$p(\mathbf{w}) \approx q(\mathbf{w}) = \frac{e^{-E(\mathbf{w}_0)}}{Z} e^{-\frac{1}{2}(\mathbf{w}-\mathbf{w}_0)^T \mathbf{A}(\mathbf{w}-\mathbf{w}_0)} = \mathcal{N}(\mathbf{w}|\mathbf{w}_0, \mathbf{A}^{-1})$$

- Approximation for Z :

$$Z = e^{-E(\mathbf{w}_0)} (2\pi)^{M/2} |\mathbf{A}|^{-1/2}$$

Back to Bayesian logistic regression

- Posterior $p(\mathbf{w}|\mathbf{t}) = \exp[-E(\mathbf{w})]/Z$ with

$$E(\mathbf{w}) = \frac{\alpha}{2} \mathbf{w}^T \mathbf{w} - \sum_n [t_n \ln y_n + (1 - t_n) \ln(1 - y_n)]$$

- $E(\mathbf{w})$ convex, single minimum, Hessian $\alpha \mathbf{I} + \Phi^T \mathbf{R} \Phi$
- Find minimum \mathbf{w}_{MAP} , call Hessian there \mathbf{S}_N^{-1}
- Then Laplace approximation for posterior is

$$p(\mathbf{w}|\mathbf{t}) \approx q(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{w}_{\text{MAP}}, \mathbf{S}_N)$$

Predictive distribution

- Use approximate posterior:

$$p(\mathcal{C}_1|\mathbf{x}, \mathbf{t}) \approx \int p(\mathcal{C}_1|\mathbf{x}, \mathbf{w})q(\mathbf{w})d\mathbf{w} = \int \sigma(\mathbf{w}^T\phi(\mathbf{x}))q(\mathbf{w})d\mathbf{w}$$

- Call $a = \mathbf{w}^T\phi(\mathbf{x})$, then a has Gaussian distribution, with

$$\mathbb{E}[a] = \int \mathbf{w}^T\phi(\mathbf{x})q(\mathbf{w})d\mathbf{w} = \mathbf{w}_{\text{MAP}}^T\phi(\mathbf{x})$$

$$\begin{aligned}\mathbb{E}[a^2] &= \int \phi(\mathbf{x})^T\mathbf{w}\mathbf{w}^T\phi(\mathbf{x})q(\mathbf{w})d\mathbf{w} \\ &= \phi(\mathbf{x})^T(\mathbf{w}_{\text{MAP}}\mathbf{w}_{\text{MAP}}^T + \mathbf{S}_N)\phi(\mathbf{x})\end{aligned}$$

- So

$$p(\mathcal{C}_1|\mathbf{x}, \mathbf{t}) \approx \int \sigma(a)\mathcal{N}(a|\mathbf{w}_{\text{MAP}}^T\phi(\mathbf{x}), \phi(\mathbf{x})^T\mathbf{S}_N\phi(\mathbf{x}))da$$

- Can be done numerically, or analytically for inverse probit

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Generative classification

- We model **joint** distribution $p(\mathbf{x}, \mathcal{C}_k)$, rather than **conditional** distribution $p(\mathcal{C}_k|\mathbf{x})$ of class labels
- Normally separate $p(\mathbf{x}, \mathcal{C}_k) = p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)$
- Class probabilities $p(\mathcal{C}_k|\boldsymbol{\pi}) = \pi_k$
- Class conditional densities e.g.

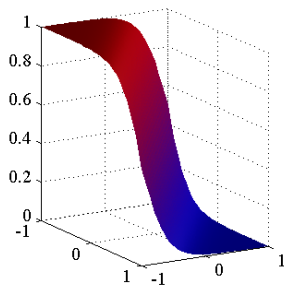
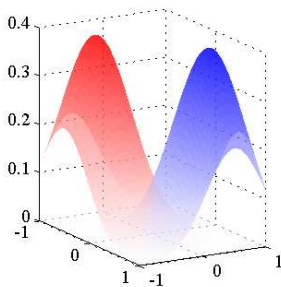
$$p(\mathbf{x}|\mathcal{C}_k, \{\boldsymbol{\mu}_j\}, \boldsymbol{\Sigma}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma})$$

- For two classes this gives

$$p(\mathcal{C}_1|\mathbf{x}) = \sigma(\mathbf{w}^T \mathbf{x} + w_0)$$

- **Linear discriminant** as before, logistic sigmoid arises naturally
- If classes have different $\boldsymbol{\Sigma}$, get **quadratic discriminant**

Illustration



Maximum likelihood inference

- Consider **two classes**, so that $\pi_1 \equiv \pi$, $\pi_2 = 1 - \pi$
- Training data: N inputs \mathbf{x}_n , N outputs $t_n \in \{0, 1\}$
- Collect into $\mathbf{X}^T = (\mathbf{x}_1, \dots, \mathbf{x}_N)$ and $\mathbf{t} = (t_1, \dots, t_N)$
- $t_n = 1$ for \mathcal{C}_1 , $t_n = 0$ for \mathcal{C}_2
- Likelihood: $p(\mathbf{x}, t = 1) = p(\mathbf{x}|\mathcal{C}_1)p(\mathcal{C}_1) = \pi\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_1, \boldsymbol{\Sigma})$
- Similarly, $p(\mathbf{x}, t = 0) = p(\mathbf{x}|\mathcal{C}_2)p(\mathcal{C}_2) = (1 - \pi)\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_2, \boldsymbol{\Sigma})$
- Overall data likelihood $p(\mathbf{t}, \mathbf{X}|\pi, \boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}) =$

$$\prod_{n=1}^N [\pi\mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_1, \boldsymbol{\Sigma})]^{t_n} [(1 - \pi)\mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_2, \boldsymbol{\Sigma})]^{1-t_n}$$

- Can be maximized in closed form

Bayesian inference

- Allow $\Sigma_1 \neq \Sigma_2$ now so each $p(\mathbf{x}|\mathcal{C}_k)$ has its own parameters
- Likelihood factorizes: $p(\mathbf{t}, \mathbf{X}|\pi, \boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}) =$

$$\prod_{n=1}^N \pi^{t_n} (1 - \pi)^{1-t_n} \prod_{n:t_n=1} \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1) \prod_{n:t_n=0} \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2)$$

- So if prior factorizes into $p(\pi)p(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1)p(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2)$, then posterior $p(\pi, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2|\mathbf{t}, \mathbf{X})$ factorizes in the same way
- Predictive distributions simplify accordingly, e.g. $p(\mathbf{x}, \mathcal{C}_1) =$

$$\int d\pi \pi p(\pi|\mathbf{t}, \mathbf{X}) \times \int d\boldsymbol{\mu}_1 d\boldsymbol{\Sigma}_1 \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1) p(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1|\mathbf{t}, \mathbf{X})$$

- Effectively, each class density models $p(\mathbf{x}|\mathcal{C}_k)$ is learnt separately from training data with class label k
- Conjugate priors $p(\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$: Gamma-Wishart