MA1972
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## SOLUTIONS to statistics questions

A5
a) $\sum_{i} x_{i}=\mathbf{2 6 1}$ Mean $=\frac{\sum_{i} x_{i}}{n}=\frac{261}{10}=\mathbf{2 6 . 1}$
b) $\sum_{i} x_{i}^{2}=6961$


$$
\sqrt{\frac{6961-6812.1}{9}}=\sqrt{\frac{148.9}{9}}=\sqrt{16.5444}=4.0675
$$

A6 $\quad E[x]=\mu \quad \operatorname{Var}[x]=\sigma^{2}$

$$
\begin{aligned}
&\left(E[n x]=n E[x]=\underline{\underline{n \mu}} \quad \operatorname{Var}[n x]=E\left[(n x)^{2}\right]-(E[n x])^{2}\right. \\
&=n^{2} E\left[(x)^{2}\right]-n^{2}(E[x])^{2} \\
&=n^{2} \operatorname{Var}[x]={\underline{\underline{n^{2}} \sigma^{2}}}^{\underline{2}}
\end{aligned}
$$

A7

$$
f(x)=\left\{\begin{array}{cc}
\frac{c}{x^{4}} & x \geq 100 \\
0 & x<100
\end{array}\right.
$$

(a). $\int_{100}^{\infty} \frac{c}{x^{4}} d x=1 \quad$ ie $c\left[\frac{x^{-3}}{-3}\right]_{100}^{\infty}=c\left[\frac{1}{3(100)^{3}}\right]=1$

$$
\therefore \quad c=3(100)^{3}=3\left(10^{6}\right)
$$

$$
\begin{align*}
& E[x]=\int_{-\infty}^{\infty} x f(x) d x=\int_{-\infty}^{100} x(0) d x+\int_{100}^{\infty} \frac{3\left(10^{6}\right)}{x^{3}} d x  \tag{b}\\
&=0+3\left(10^{6}\right)\left[\frac{1}{-2 x^{2}}\right]_{100}^{\infty}=0+3\left(10^{6}\right)\left[0+\frac{1}{2\left(100^{2}\right)}\right] \\
&=\frac{3\left(10^{6}\right)}{2\left(10^{4}\right)}=\underline{\underline{150}} \\
& E\left[x^{2}\right]=\int_{-\infty}^{\infty} x^{2} f(x) d x=0+\int_{100}^{\infty} \frac{3\left(10^{6}\right)}{x^{2}} d x=3\left(10^{6}\right)\left[\frac{1}{-x}\right]_{100}^{\infty} \\
&=\frac{3\left(10^{6}\right)}{10^{2}}=\underline{3\left(10^{4}\right)}
\end{align*}
$$

$$
\operatorname{Var}[x]=E\left[x^{2}\right]-(E[x])^{2}=3\left(10^{4}\right)-\left[(1.5)\left(10^{2}\right)\right]^{2}=10^{4}\left[3-(1.5)^{2}\right]=\underline{\underline{7500}}
$$

A8

$$
n=100 \quad p=\frac{59}{100}=0.59 \quad \therefore q=0.41
$$

Since $n>30$ the proportion of students in debt is $\approx$ normal
The $95 \%$ CI for $\pi$ is $\quad p \pm z \sqrt{\frac{p(q)}{n}}=0.59 \pm 1.96 \sqrt{\frac{0.59(0.41)}{100}}$

$$
=0.59 \pm 0.096
$$

$$
\underline{\underline{i e}[0.494,0.686]}
$$

A9

$$
\sigma \text { is known }=16 \quad \mathrm{n}=15 \quad \mathrm{~s}=13.03
$$

$H_{0}$ : Urban area standard deviation in IQ is the same as the whole country $\sigma=16$
$H_{1}$ : Urban area standard deviation in IQ is not the same as whole country $\sigma \neq 16$
At 5\% level - Critical values are from $\chi^{2}$ distribution $\quad \chi_{L 2.5 \%}^{2}=\underline{\underline{5.629}}$

$$
\chi_{U 2.5 \%}^{2}=\underline{\underline{26.12}}
$$

$\chi^{2}=\frac{(n-1) s^{2}}{\sigma^{2}}=\frac{14(169.8)}{256}=\underline{\underline{9.286}}$
$\therefore$ Accept $\mathrm{H}_{0}$


There is no evidence to suggest that the variability of IQ's in urban areas is different from the whole country.
a) State hypotheses
$\mathrm{H}_{0}$ : The new lacquer provides on average the same protection against rust $\mu_{1}=\mu_{2}$
$H_{1}$ : The new lacquer provides added protection against rust $\mu_{2}>\mu_{1}$ (one tail)
State level of significance e.g. let $\alpha=5 \%$ then critical value $=1.64$
Test Statistic $\mathbf{Z}=\frac{\overline{x_{1}}-\overline{x_{2}}}{\sqrt{\frac{\mathrm{~s}_{1}^{2}}{\mathrm{n}_{1}}+\frac{\mathrm{s}_{2}^{2}}{\mathrm{n}_{2}}}}=\frac{730-480}{\sqrt{\frac{80^{2}}{256}+\frac{100^{2}}{50}}}=\frac{250}{\sqrt{25+200}}$

$$
=\frac{250}{\sqrt{225}}=\frac{250}{15}=\underline{\mathbf{1 6 . 6 6 7}}
$$

Since $\mathrm{Z}>1.64$ reject $\mathrm{H}_{0}$ and accept $\mathrm{H}_{1}$, there is a significant result.
CONCLUSION: there is evidence that the new lacquer provides added protection against rust.
b)

|  | OBSERVED |  |  |  |  | EXPECTED |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :--- | :--- | :--- | :---: | :---: | ---: | ---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Firm | No | No | Total |  | Firm | No | No | Total |  |  |  |  |  |  |  |  |  |
| Satisfied |  |  |  |  |  |  |  |  |  |  | Not |  |  |  | Satisfied | Not |  |
| Customers |  |  | Satisfied |  |  | Customers |  |  | Satisfied |  |  |  |  |  |  |  |  |
| A | 576 | 24 | 600 |  | A | 565 | 35 | 600 |  |  |  |  |  |  |  |  |  |
| B | 558 | 42 | 600 |  | B | 565 | 35 | 600 |  |  |  |  |  |  |  |  |  |
| C | 580 | 20 | 600 |  | C | 565 | 35 | 600 |  |  |  |  |  |  |  |  |  |
| D | 546 | 54 | 600 |  | D | 565 | 35 | 600 |  |  |  |  |  |  |  |  |  |
|  | 2260 | 140 | 2400 |  |  | 2260 | 140 | 2400 |  |  |  |  |  |  |  |  |  |

We have a 4 X 2 contingency table problem, and can test the hypotheses,
$\mathrm{H}_{0}$ : The coach firms are the same service
$\mathrm{H}_{1}$ : There is a difference between the coach firms.
If $\alpha=5 \% \quad \chi_{3 \times 1}{ }^{2}=7.815$

$$
\begin{aligned}
\chi^{2}= & \frac{\sum_{i}\left(O_{i}-E_{i}\right)^{2}}{E_{i}}=1 / 565\left(11^{2}+7^{2}+15^{2}+19^{2}\right)+1 / 35\left(11^{2}+7^{2}+15^{2}+19^{2}\right) \\
= & 756 / 565+756 / 35=1.3380+21.6=\mathbf{2 2 . 9 3 8}
\end{aligned}
$$

So we reject Ho, there is strong evidence that the firms have different levels of satisfaction amongst its customers.
a) British
$\mathrm{n}_{1}=20 \quad \overline{\mathrm{X}}_{1}=3.7$
$\mathrm{S}_{1}=0.6$
American $\quad \mathrm{n}_{2}=15 \quad \overline{\mathrm{X}}_{2}=4.2 \quad \mathrm{~S}_{2}=0.9$

State hypotheses
$\mathrm{H}_{0}$ : British \& American companies spend on average the same on R\&D $\mu_{1}=\mu_{2}$
$H_{1}$ :British companies do not spend as much on R\&D as American companies $\mu_{2}>\mu_{1}$ (one tail)

We will assume that the distribution of expenditure is normal and the population variances are unknown but equal. Since the sample sizes are small we will use the $t$ test with d.f. $=n_{1}+n_{2}-2=33$, and pooled sample standard deviation, $S$, where
$\mathrm{S}^{2}=\frac{\left(\mathrm{n}_{1}-1\right) \mathrm{S}_{1}^{2}+\left(\mathrm{n}_{2}-1\right) \mathrm{S}_{2}^{2}}{\left(\mathrm{n}_{1}-1\right)+\left(\mathrm{n}_{2}-1\right)}=\frac{19(0.6)^{2}+14(0.9)^{2}}{20+15-2}=\frac{19(0.36)+14(0.81)}{33}=\mathbf{0 . 5 5}$
Then the critical value of t for $\alpha=5 \%$ with d.f. $=\left(\mathrm{n}_{1}+\mathrm{n}_{2}-2\right)=33$ is $\mathrm{t}_{05}=\mathbf{- 1 . 6 9 2}$.
So we reject H0 if T <-1.692
Since $S^{2}=0.55$. $\mathbf{S}=\mathbf{0 . 7 4 1 6}$
The test statistic is $T=\frac{\bar{X}_{1}-\bar{X}_{2}}{S \sqrt{\frac{1}{n_{1}}+\frac{1}{n_{2}}}}=\frac{3.7-4.2}{0.7416 \sqrt{\frac{1}{20}+\frac{1}{15}}}=\mathbf{- 1 . 9 7}$

$$
-1.97<\mathrm{t}_{.05}=-1.692, \text { so we reject Ho }
$$

We conclude that there is evidence to support the view that British companies spend less on R \& D than American companies.
b (i) We wish to test the hypotheses
Ho:The variances in expenditure are the same $\quad \sigma_{1}^{2}=\sigma_{2}^{2}$
H 1 : The variances in expenditure are not the same $\sigma_{1}{ }^{2} \neq \sigma_{2}^{2}$
At $\alpha=5 \%$ significance level. The right-tailed critical value of F with $v_{2}=(15-1)=$ 14 and $v_{1}=(20-1)=19 \quad \mathrm{U}_{0.025}=2.65$.

The test statistic is $\mathrm{F}=\frac{s_{2}^{2}}{s_{1}^{2}}=\frac{(0.9)^{2}}{(0.6)^{2}}=\mathbf{2 . 2 5}$
We accept Ho: $\sigma_{1}^{2}=\sigma_{2}^{2}$.
Thus, there is no evidence to suggest unequal variances.
(ii) There is no evidence against equal variances, therefore the $t$ test is valid.

