

Cambridge International Examinations Cambridge Pre-U Certificate

MATHEMATICS (STATISTICS WITH PURE MATHEMATICS) (SHORT COURSE)

For Examination from 2016

SPECIMEN MARK SCHEME

Paper 2 Statistics

2 hours

1347/02

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MAXIMUM MARK: 80

The syllabus is approved for use in England, Wales and Northern Ireland as a Cambridge International Level 3 Pre-U Certificate.

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Mark Scheme Notes

Marks are of the following three types:

- M Method mark, awarded for a valid method applied to the problem. Method marks are not lost for numerical errors, algebraic slips or errors in units. However it is not usually sufficient for a candidate just to indicate an intention of using some method or just to quote a formula; the formula or idea must be applied to the specific problem in hand, e.g. by substituting the relevant quantities into the formula. Correct application of a formula without the formula being quoted obviously earns the M mark and in some cases an M mark can be implied from a correct answer.
- A Accuracy mark, awarded for a correct answer or intermediate step correctly obtained. Accuracy marks cannot be given unless the associated method mark is earned (or implied).
- B Mark for a correct result or statement independent of method marks.

The following abbreviations may be used in a mark scheme:

- AG Answer Given on the question paper (so extra checking is needed to ensure that the detailed working leading to the result is valid)
- CAO Correct Answer Only (emphasising that no "follow through" from a previous error is allowed)
- aef Any equivalent form
- art Answers rounding to
- cwo Correct working only (emphasising that there must be no incorrect working in the solution)
- ft Follow through from previous error is allowed
- o.e. Or equivalent

| 1 | (i) | $S_{xx} = 1939552 - \frac{(4412)^2}{13} = 442187$ (to nearest integer) | B1 |
|---|-------|--|----------|
| | | $S_{yy} = 605147 - \frac{(2387)^2}{13} = 166857$ (to nearest integer) | B1 |
| | | $S_{xy} = 1074848 - \frac{4412 \times 2387}{13} = 264737$ (to nearest integer) | B1 |
| | | $r = \frac{264737}{\sqrt{442187 \times 166857}} = 0.975 \ (0.9746)$ | |
| | | Calculating r from their S_{xx} , S_{yy} and S_{xy} | M1 |
| | | (numerical working or <i>their r</i> value correct to 3 sf or better) r is near 1, so a good fit to an upward sloping line | A1 |
| | | Drawing a valid conclusion (confirming that a linear fit is appropriate, as stated in question) | |
| | (ii) | $b = \frac{264737}{442197} = 0.599 \ (0.5987)$ | M1 |
| | | Calculating b from their S_{xx} and S_{xy} | |
| | | $a = \frac{2387}{12} - 0.5987 \times \frac{4412}{12}$ | M1 |
| | | Calculating <i>a</i> from $\sum x$, $\sum y$ and <i>their b</i> | |
| | | $= 183.6 - 0.5987 \times 339.4 = -19.6$ | |
| | | y = 0.599x - 19.6 Line correct with coefficients to 3sf or better | A1 |
| | | $x = 2203 \Rightarrow \hat{y} = 1300$ (From their line (± 2)) | B1 |
| | (iii) | Extrapolation beyond range of data | B1 |
| | | Any valid objection | B1 |
| | | Small sample / only based on one sample Sampling method not known / not random sampling | |
| | | London is not typical / London 'is different' | |
| 2 | (i) | Median = 30 mpg CAO | B1 |
| | | Quartiles = 34 mpg and 23 mpg (Accept 33 to 35 and 20 to 24) IQR = 11 mpg (<i>their</i> IQR calculated) | B1 M1 |
| | | Outliers have mpg < 6.5 or > 50.5 | A1 |
| | | \Rightarrow Toyota Prius | B1 |
| | (ii) | Using median and quartile values appropriately to deduce non-normal | B1 |
| | | e.g. The difference between 23 and 30 is much greater than the difference between 30 and 34; this suggests that the distribution is not symmetric. | |

| | (iii) | 1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 | |
|---|-------|--|----------|
| | | $\frac{1 \ 2 \ 12 \ 3 \ 4 \ 5 \ 6 \ 7 \ 8 \ 11 \ 9 \ 10 \ 13 \ 14 \ 15}{0 \ 0 \ -9 \ 1 \ 1 \ 1 \ 1 \ 1 \ 1 \ -1 \ 2 \ 2 \ 0 \ 0 \ 0}$ | |
| | | Substantially correct calculation of d or $ d $ or d^2 for the ranks | M1 |
| | | | . 1 |
| | | $\Sigma d^2 = 96$ | A1 |
| | | $r_s = 1 - \frac{6 \times 96}{15 \times 224} = 1 - 0.17143 = 0.82857$ | |
| | | Substantially correct calculation of d or $ d $ or d^2 for the ranks | M1 |
| | | = 0.829 (3 sf) | A1 |
| 3 | (i) | Independence between children, (random sample) class are typical of population in respect of left-handedness (Independence Probability 13% (constant probability)) | B1 B1 |
| | (ii) | X = number of left-handers $X \sim B(20, 0.13)$ | |
| | | 13% of 20 = 2.6, so want $P(X \le 2)$ | B1 |
| | | $(0.87)^{20} + 20(0.13)(0.87)^{19} + 190(0.13)^2(0.87)^{18}$ | M1 |
| | | Calculating a probability in B(20, 0.13) (At least) three correct probabilities added | A1 |
| | | = 0.061714 + 0.18443 + 0.26181 | |
| | | = 0.50795 = 0.508 to 3sf | A1 |
| | (iii) | $X \sim B(20, p)$ $p = P(left-hander) (may be implied)$ | |
| | | $H_0: p = 0.13$ | B1 |
| | | $H_1: p > 0.13$ | B1 |
| | | $\alpha = 5\%$ one-tailed test (Omission of <i>p</i> only penalised once. May imply level of test and one-tailed) | |
| | | Assuming $H_0, X \sim B(20, 0.13)$ | |
| | | $P(X \ge 7) = 1 - 0.9897 = 0.0103$ | M1 |
| | | or cv = 6 0.0103 < 5% or 7 > 6 | M1 |
| | | Reject H_0 | A1 |
| | | Evidence supports claim, significantly more of the most recent twenty presidents were left-handed than would be expected by chance. (must be in context) | B1 |
| | (iv) | Any valid reason, either from context or addressing statistical variation | B1 |
| | () | e.g. Schools trained pupils to write with their right hand in the past | |
| | | Left-handedness was not recorded accurately in the past | |
| | | Not random samples, could be due to sample variation | |

| r | | | |
|---|-------|---|------|
| 4 | (i) | $P(Z > z) = 0.01 \Longrightarrow z = 2.326$ | |
| | | $P(Z < z) = 0.25 \Rightarrow z = -0.674$ (both values required for mark) | B1 |
| | | Substantially correct method | M1 |
| | | $2.326 = \frac{120 - \mu}{\sigma} \Rightarrow 120 - \mu = 2.326\sigma$ | |
| | | $-0.674 = \frac{84 - \mu}{\sigma} \Longrightarrow 84 - \mu = -0.674\sigma$ | |
| | | 0 | A1 |
| | | Both correct for <i>their z</i> -values, one of which is positive and one negative | |
| | | $\Rightarrow \mu = 92.1 \text{ CAO}$ | B1 |
| | | $\sigma = 12 \text{ CAO}$ | B1 |
| | | | |
| | (ii) | H ₀ : samples come from same populations | |
| | | H ₁ : S tend to have larger increases than N | B1 |
| | | (S have smaller rank values than N) | |
| | | Appropriate statement of hypotheses | |
| | | One-tailed test, $\alpha = 5\%$ | |
| | | Rank sum for $S = 1 + 3 + 4 + 5 + 6 + 8$ | M1 |
| | | $\Rightarrow W = 27$ | A1 |
| | | $m = 6$ $n = 10 \Rightarrow$ critical value for $W = 35$ | B1 |
| | | Deject II | |
| | | Reject H ₀ (Correct conclusion, in context) | B1 |
| | | At the 5% level the data support the claim that the increases are greater for the | |
| | | smokers than for the non-smokers | |
| | | | |
| | (iii) | For the smokers, $\sum x = 708 \Rightarrow \overline{x} = 118$ | D1 |
| | | Estimate $\hat{\mu}_s = 118$ CAO | B1 |
| | | $\sum x^2 = 83864 \implies S_{xx} = 320 \implies s^2 = 64$ | M1 |
| | | Sight of one of 83864, 320, 64, 8, 53.3 or 7.30 | 1111 |
| | | Estimate $\hat{\sigma}_s^2 = 64$ CAO | A1 |
| | | | |
| | (iv) | $\overline{X} \sim N(\mu_s, \frac{\sigma_s^2}{n})$ where $\widehat{\sigma_s} = 8$ and $n = 6$ | |
| | | Critical values in t(5) are ± 2.571 | |
| | | Using t tables to find 2.571 or 2.447 | B1 |
| | | Confidence interval is $118 \pm 2.571 \times \frac{8}{\sqrt{6}}$ | M1 |
| | | Correct method for <i>their</i> "t" value and <i>their</i> \overline{x} , $\hat{\sigma}$ | |
| | | $= 118 \pm 8.4 = [109.6, 126.4]$ (ft <i>their</i> values from part (iii)) | A1 |
| | | | |

| 5 | (i) | $x = 47 \rightarrow z = 0.667$; $x = 51 \rightarrow z = 2.0$ (z values may be implied) | B1 |
|---|----------|--|----|
| | | $P(47 < X \le 51) = 0.9772 - 0.7477 $ (may be implied) | M1 |
| | | Expected frequency = $0.2295 \times 100 = 22.95$ AG | A1 |
| | (ii) | Merge classes in tails to make expected frequencies at least 5 | B1 |
| | | Weight<43 $43-45$ $45-47$ >47Observed frequency32243014Expected frequency25.2324.7724.7725.23 $(O-E)^2/E$ 1.820.021.105.00 | |
| | | Substantially correct calculation of X^2 (with or without merging) | M1 |
| | | X^2 calculated = art 7.94 CAO | A1 |
| | | H ₀ : N(45, 9) distribution H ₁ : some other distribution | B1 |
| | | From tables, critical value = 7.815 CAO | B1 |
| | | Reject H_0 Data is not consistent with a N(45, 9) distribution | B1 |
| | (iii) | v = n - 1 = 4 - 1 = 3 | B1 |
| | | No need to reduce degrees of freedom for parameters as not estimated from sample data | |
| | (iv) (a) | Variance cannot be estimated, midpoints cannot be found for first and last classes since boundaries are not known | B1 |
| | (b) | Sign test or binomial test or equivalent (e.g. test proportion that are below 45) | B1 |
| | | H_0 : median = 45 $\alpha = 5\%$ H_1 : median $\neq 45$ two-tailed test | |
| | | $Y =$ number of chicks with weight ≤ 45 g Assuming H ₀ , $Y \sim B(100, 0.5)$ | M1 |
| | | Approximate by N(50, 25) Critical values are $50 \pm 1.96 \times 5 = 50 \pm 9.8$ = [40.2, 59.8] | A1 |
| | | Observed $y = 56$ (or 44 above) | |
| | | Accept H_0 Data are consistent with a distribution with median = 45. No evidence that median is not 45 | B1 |

| 6 | (i) | $X \sim N(10, 9)$ approx | |
|---|-------|--|------|
| | | Correct mean | M1 |
| | | Correct variance | A1 |
| | | <i>their</i> mean $+$ 1.645 \times <i>their</i> sd (with or without continuity correction) | M1 |
| | | Critical value = $10 + 1.645 \times 3 + 0.5$ | |
| | | = 14.935 + 0.5 | |
| | | = 15.435 | |
| | | Critical value = 16 | A1 |
| | | | |
| | | If the number observed is 15 or fewer, accept H_0 and conclude that p may be 0.10 | B1 |
| | | If number observed is 16 or more, reject H_0 and conclude that p is probably greater | |
| | | than 0.10 | |
| | | P(Type I error) | B1 |
| | | = $P(reject H_0 when it is true)$ | |
| | | (ft their integer cv of 16 or 15) | |
| | | $= P(X \ge 16)$ in B(100, 0.10) | M1 |
| | | $= P(X \ge 15.5)$ in N(10, 9) approx. | |
| | | $= P(Z \ge (15.5 - 10)/3) = P(Z \ge 1.833) = 1 - 0.9666$ | A1 |
| | | = 0.0334 | |
| | | | |
| | (ii) | P(Type II error) = P(accept H ₀ when it is false) = P($X \le 15$) in B(100, 0.20) | B1 |
| | | B(100, 0.20) or $N(20, 16)$ used | M1 |
| | | | |
| | | $= P(X \le 15.5)$ in N(20, 16) approx. | |
| | | ft <i>their</i> integer cv of 16 or 15 | |
| | | $= P(Z \le (15.5 - 20)/4) = P(Z \le -1.125) = 1 - 0.8696$ | A1 |
| | | = 0.1304 | |
| | | 0.1504 | |
| | (iii) | $P \sim N(0.14, \frac{0.14 \times 0.86}{100})$ | B1 |
| | () | | 21 |
| | | = N(0.14, 0.001204) approx. | |
| | | Correct method for <i>their</i> distribution | N (1 |
| | | 95% CI = $0.14 + 1.96\sqrt{0.001204}$ | M1 |
| | | — | A 1 |
| | | $= 0.14 \pm 0.068 = [0.072, 0.208]$ Correct interval, in any form, with or | A1 |
| | | without an attempt at continuity 0.10 and 0.20 are both in this interval | D1 |
| | | 0.10 and 0.20 are both in this interval | B1 |

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