

# ADVANCED GCE MATHEMATICS (MEI)

4769/01

Statistics 4

FRIDAY 6 JUNE 2008 Afternoon

Time: 1 hour 30 minutes

Additional materials: Answer Booklet (8 pages)

Graph paper

MEI Examination Formulae and Tables (MF2)

## **INSTRUCTIONS TO CANDIDATES**

- Write your name, centre number and candidate number in the spaces provided on the answer booklet.
- Read each question carefully and make sure you know what you have to do before starting your answer.
- Answer any three questions.
- You are permitted to use a graphical calculator in this paper.
- Final answers should be given to a degree of accuracy appropriate to the context.

#### INFORMATION FOR CANDIDATES

- The number of marks is given in brackets [] at the end of each question or part question.
- The total number of marks for this paper is 72.
- You are advised that an answer may receive **no marks** unless you show sufficient detail of the working to indicate that a correct method is being used.

This document consists of 4 printed pages.

## Option 1: Estimation

1 The random variable X has the Poisson distribution with parameter  $\theta$  so that its probability function is

$$P(X = x) = \frac{e^{-\theta} \theta^x}{x!}, \qquad x = 0, 1, 2, \dots,$$

where  $\theta$  ( $\theta > 0$ ) is unknown. A random sample of *n* observations from *X* is denoted by  $X_1, X_2, \ldots, X_n$ .

(i) Find 
$$\hat{\theta}$$
, the maximum likelihood estimator of  $\theta$ . [9]

The value of P(X = 0) is denoted by  $\lambda$ .

- (ii) Write down an expression for  $\lambda$  in terms of  $\theta$ .
- (iii) Let R denote the number of observations in the sample with value zero. By considering the binomial distribution with parameters n and  $e^{-\theta}$ , write down E(R) and Var(R). Deduce that the observed *proportion* of observations in the sample with value zero, denoted by  $\tilde{\lambda}$ , is an unbiased estimator of  $\lambda$  with variance  $\frac{e^{-\theta}(1-e^{-\theta})}{n}$ .
- (iv) In large samples, the variance of the maximum likelihood estimator of  $\lambda$  may be taken as  $\frac{\theta e^{-2\theta}}{n}$ . Use this and the appropriate result from part (iii) to show that the relative efficiency of  $\tilde{\lambda}$  with respect to the maximum likelihood estimator is  $\frac{\theta}{e^{\theta}-1}$ . Show that this expression is always less than 1. Show also that it is near 1 if  $\theta$  is small and near 0 if  $\theta$  is large. [7]

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#### **Option 2: Generating Functions**

- Independent trials, on each of which the probability of a 'success' is p (0 ), are being carried out. The random variable <math>X counts the number of trials up to and including that on which the first success is obtained. The random variable Y counts the number of trials up to and including that on which the nth success is obtained.
  - (i) Write down an expression for P(X = x) for x = 1, 2, ... Show that the probability generating function of X is

$$G(t) = pt(1 - qt)^{-1}$$

where q = 1 - p, and hence that the mean and variance of X are

$$\mu = \frac{1}{p}$$
 and  $\sigma^2 = \frac{q}{p^2}$ 

respectively. [11]

(ii) Explain why the random variable Y can be written as

$$Y = X_1 + X_2 + \ldots + X_n$$

where the  $X_i$  are independent random variables each distributed as X. Hence write down the probability generating function, the mean and the variance of Y. [5]

- (iii) State an approximation to the distribution of Y for large n. [1]
- (iv) The aeroplane used on a certain flight seats 140 passengers. The airline seeks to fill the plane, but its experience is that not all the passengers who buy tickets will turn up for the flight. It uses the random variable Y to model the situation, with p = 0.8 as the probability that a passenger turns up. Find the probability that it needs to sell at least 160 tickets to get 140 passengers who turn up.

Suggest a reason why the model might not be appropriate. [7]

# Option 3: Inference

3 (i) Explain the meaning of the following terms in the context of hypothesis testing: Type I error, Type II error, operating characteristic. [6]

A machine fills salt containers that will be sold in shops. The containers are supposed to contain 750 g of salt. The machine operates in such a way that the amount of salt delivered to each container is a Normally distributed random variable with standard deviation 20 g. The machine should be calibrated in such a way that the mean amount delivered,  $\mu$ , is 750 g.

Each hour, a random sample of 9 containers is taken from the previous hour's output and the sample mean amount of salt is determined. If this is between 735 g and 765 g, the previous hour's output is accepted. If not, the previous hour's output is rejected and the machine is recalibrated.

- (ii) Find the probability of rejecting the previous hour's output if the machine is properly calibrated. Comment on your result. [6]
- (iii) Find the probability of accepting the previous hour's output if  $\mu = 725$  g. Comment on your result.
- (iv) Obtain an expression for the operating characteristic of this testing procedure in terms of the cumulative distribution function  $\Phi(z)$  of the standard Normal distribution. Evaluate the operating characteristic for the following values (in g) of  $\mu$ : 720, 730, 740, 750, 760, 770, 780. [6]

# Option 4: Design and Analysis of Experiments

- 4 (i) State the usual model, including the accompanying distributional assumptions, for the one-way analysis of variance. Interpret the terms in the model. [9]
  - (ii) An examinations authority is considering using an external contractor for the typesetting and printing of its examination papers. Four contractors are being investigated. A random sample of 20 examination papers over the entire range covered by the authority is selected and 5 are allocated at random to each contractor for preparation. The authority carefully checks the printed papers for errors and assigns a score to each to indicate the overall quality (higher scores represent better quality). The scores are as follows.

Contractor A	Contractor B	Contractor C	Contractor D
41	54	56	41
49	45	45	36
50	50	54	46
44	50	50	38
56	47	49	35

[The sum of these data items is 936 and the sum of their squares is 44 544.]

Construct the usual one-way analysis of variance table. Carry out the appropriate test, using a 5% significance level. Report briefly on your conclusions. [12]

(iii) The authority thinks that there might be differences in the ways the contractors cope with the preparation of examination papers in different subject areas. For this purpose, the subject areas are broadly divided into mathematics, sciences, languages, humanities, and others. The authority wishes to design a further investigation, ensuring that each of these subject areas is covered by each contractor. Name the experimental design that should be used and describe briefly the layout of the investigation.

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# 4769 Statistics 4

_				
Q1 (i)				
(1)				
	, $e^{- heta} heta^{x_1}$ $e^{- heta} heta^{x_n}$ $\left[ e^{-n heta} heta^{\sum x_i}  ight]$	M1	product form	
	$L = \frac{e^{-\theta}\theta^{x_1}}{x_1!} \dots \frac{e^{-\theta}\theta^{x_n}}{x_n!} = \frac{e^{-n\theta}\theta^{\sum x_i}}{x_1!x_2! \dots x_n!}$	A1	fully correct	
	$ \ln L = const - n\theta + \sum x_i \ln \theta $	M1 A1		
	$d \ln I \qquad \sum r$			
	$\frac{d \ln L}{d\theta} = -n + \frac{\sum x_i}{\theta} = 0$	M1 A1		
	$\Rightarrow \hat{\theta} = \frac{\sum x_i}{(= \overline{x})}$	AI		
	$\Rightarrow \theta = \frac{\sum_{i=1}^{n} (=\bar{x})}{n}$	A1	CAO	
	Check this is a maximum	M1		
	e.g. $\frac{d^2 \ln L}{d a^2} = -\frac{\sum x_i}{a^2} < 0$			
	e.g. $\frac{d\theta^2}{d\theta^2} = -\frac{2}{\theta^2} < 0$			
		A1		
				9
(ii)	$\lambda = P(X = 0) = e^{-\theta}$	B1		1
(iii)	We have $R \sim \mathrm{B}(n,e^{-\theta})$ ,	M1		
	so $E(R) = ne^{-\theta}$	В1		
	$Var(R) = ne^{-\theta} (1 - e^{-\theta})$	В1		
	$\sim R$			
	$\widetilde{\lambda} = \frac{R}{n}$	M1		
		A1		
	$\therefore E(\widetilde{\lambda}) = e^{-\theta}$	A1		
	i.e. unbiased $e^{-\theta}(1, e^{-\theta})$	A 4	BEWARE PRINTED	
	$\operatorname{Var}(\widetilde{\lambda}) = \frac{e^{-\theta}(1 - e^{-\theta})}{n}$	A1	ANSWER	_
	rı			7

(iv)	Relative efficiency of $\widetilde{\lambda}$ wrt ML est $= \frac{\text{Var}(\text{ML Est})}{\text{Var}(\widetilde{\lambda})}$ $= \frac{\theta e^{-2\theta}}{n} \cdot \frac{n}{e^{-\theta} (1 - e^{-\theta})} = \frac{\theta}{e^{\theta} - 1}$	M1 M1 A1	any attempt to compare variances if correct BEWARE PRINTED ANSWER	
	Eg:- Expression is $\frac{\theta}{\theta + \frac{\theta^2}{2!} + \dots}$	M1		
	always < 1 and this is $\approx$ 1 if $\theta$ is small $\approx$ 0 if $\theta$ is large	E1 E1 E1	Allow statement that $\frac{\theta}{e^{\theta}-1} \to 0 \text{ as } \theta \to \infty$	7

02		I	<u> </u>	
(i)	D(V ) x-1	B1	FT into pgf only	
(1)	$P(X=x) = q^{x-1}p$	51	1 1 litto pgi only	
	Pgf $G(t) = E(t^{x}) = \sum_{x=1}^{\infty} pt^{x} q^{x-1}$	M1		
	$= pt(1 + qt + q^{2}t^{2} +)$ $= \underline{pt(1 - qt)^{-1}}$	A1 A1	BEWARE PRINTED ANSWER	
	$\mu = G'(1)$ $\sigma^2 = G''(1) + \mu - \mu^2$	M1	[consideration of  qt  < 1 not required]	
	$G'(t) = pt(-1)(1-qt)^{-2}(-q) + p(1-qt)^{-1}$		for attempt to find G'(t) and/or G"(t)	
	$= pqt(1-qt)^{-2} + p(1-qt)^{-1}$	A1		
	$\therefore G'(1) = pq(1-q)^{-2} + p(1-q)^{-1} = \frac{q}{p} + 1 = \frac{1}{\underline{p}}$ $G''(t) = pqt(-2)(1-qt)^{-3}(-q) + pq(1-qt)^{-2} +$	A1	BEWARE PRINTED ANSWER	
	$p(-1)(1-qt)^{-2}(-q)$	A1		
	$\therefore G''(1) = 2pq^{2}(1-q)^{-3} + pq(1-q)^{-2} + pq(1-q)^{-2}$ $= \frac{2q^{2}}{p^{2}} + \frac{2q}{p}$			
	P P	A1		
	$\therefore \sigma^2 = \frac{2q^2}{p^2} + \frac{2q}{p} + \frac{1}{p} - \frac{1}{p^2} = \frac{2q^2 + 2pq + p - 1}{p^2}$	M1	For inserting their values	
	$= \frac{q}{p^2} (2q + 2p - 1) = \frac{q}{\underline{p^2}}$	A1	BEWARE PRINTED ANSWER	
				11

(ii)	$X_1$ =number of trials to first success			
,	$X_2$ = " " " next " $X_2$ + $X_2$ + $X_2$ + $X_2$ + $X_2$ + $X_3$			
		E1		
	. = total no of trials	E1		
	to the <i>n</i> th success	- '		
	X <sub>n</sub> = " " " " nth "			
	Λη- <i>I</i> III			
	$\therefore \operatorname{pgf} \operatorname{of} Y = (\operatorname{pgf} \operatorname{of} X)^n = p^n t^n (1 - qt)^{-n}$	1		
	<del></del>			
	$n = nn = \frac{n}{n}$			
	$\mu_Y - n\mu_X - p$	1		
	$\mu_{Y} = n\mu_{X} = \frac{n}{\underline{p}}$			
	$\sigma_{\rm Y}^2 = n\sigma_{\rm X}^2 = \frac{nq}{2}$			
	$\sigma_{Y} = n\sigma_{X} = \frac{1}{n^{2}}$			
	<u>p</u>	4		_
/:::\	2	1		5
(iii)	N(candidate's $\mu_{\scriptscriptstyle Y}$ , candidate's $\sigma_{\scriptscriptstyle Y}^2$ )	1		1
(iv)	Y = no of tickets to be sold ~ random variable as			
(.*)	in (ii) with $n = 140$ and $p = 0.8$	E1		
	•	- '		
	~ Approx N $\left(\frac{140}{0.8} = 175, \frac{140 \times 0.2}{(0.8)^2} = 43.75\right)$	4		
	$0.8$ $(0.8)^2$	1	Do not oward if attraction	
	$P(Y \ge 160) \approx P(N(175,43.75) > 159\frac{1}{2})$		Do not award if cty corr	
	$1(1 \ge 100) \sim 1(1(1/3,43.73) > 139 \frac{1}{2})$	M1	absent or wrong, but FT if	
			160 used →	
	= P(N(0,1)>-2.343)		-2.268, 0.9884	
	= 0.9905	A1		
	- 0.3300	A1		
	For any consible discussion in contact (as groups	′ ` `	CAO	
	For any sensible discussion in context (eg groups	E1	0,10	
	of passengers $\Rightarrow$ not indep.)			7
-00	N 1 5 11 N/ 5-703 25023	E1		7
Q3	$X = \text{amount of salt} \sim N(\mu[750], \sigma^2[20^2])$			
	Sample of <i>n</i> =9			
(i)	Type I error: rejecting null hypothesis	B1	Allow B1 for	
( )	when it is true.	В1	P(rej H₀ when true)	
	Wien to a do.	-	(i oj i io mien ado)	
	Type II error: accepting null hypothesis	В1	Allow B1 for	
	when it is false.	B1	P(acc H₀ when false)	
	00.0	٦.		
	OC: P (accepting null hypothesis	B1	[ P(type II error   the true	
	as a function of the parameter under	B1	value of the parameter)	6
	investigation)	<u></u>	scores B1+B1]	
(ii)	Reject if $\overline{x} < 735 \text{ or } \overline{x} > 765$			
` ′		M1	Might be implicit	
	$\alpha = P(\overline{X} < 735 \text{ or } \overline{X} > 765   \overline{X} \sim N(750, \frac{20^2}{9}))$		5	
	9 "			
	-P(7 < (735-750)3 - 2.25	A1		
	$= P(Z < \frac{(735 - 750)3}{20} = -2.25$	Α1		
	_ v			
	or $Z > \frac{(765 - 750)3}{20} = 2.25$	A1		
	20			
	$= 2(1-0.9878) = 2 \times 0.0122 = 0.0244$	A1	CAO	
	_(:::::::::::::::::::::::::::::::::::::			
	This is the probability of rejecting good output			
	This is the probability of rejecting good output	E1		
	and unnecessarily re-calibrating the machine –		Accept any sensible	
	seems small	E1	Accept any sensible	6
	[but not very small?]		comments	6
		<u> </u>		
		_		_

(iii)	Accept if $735 < x < 765$ , and now $\mu = 725$ .	M1	might be implicit	
	$\beta = P(735 < \overline{X} < 765 \mid \overline{X} \sim N(725, \frac{20^2}{9}))$			
	= P(1.5	A1		
	< Z< 6)	A1		
	$= 1 - 0.9332 = \underline{0.0668}$	A1	CAO	
			If upper limit 765 not considered, maximum 2 of	
			these 4 marks. If $\Phi(6)$ not	
	This is the probability of accepting output and	l	considered, maximum 3	_
	carrying on when in fact $\mu$ has slipped to 725 –	E1	out of 4.	6
	small[-ish?]	E	accept sensible comments	
(iv)	OC = P( $735 < \overline{X} < 765 \mid \overline{X} \sim N(\mu, \frac{20^2}{9})$ )	M1		
	/ /			
	$= \Phi  \left(\frac{(765 - \mu)^3}{20}\right) - \Phi  \left(\frac{(735 - \mu)^3}{20}\right)$			
	( 20 )			
	" Φ - Φ"			
	Ψ Ψ	M1	both correct	
		A1	Dou't correct	
	$\mu$ =720: $\Phi$ (6.75) – $\Phi$ (2.25)=1– 0.9878 =0.0122 730: 5.25 0.75 =1– 0.7734 =0.2266			
	740: 3.75		if any two correct	
		1	In any two correct	
	750: similarly or by write-down from part (ii)	1		
	[FT]: 0.9756	'		
	760, 770, 780 by symmetry			
	[FT]: 0.7734, 0.2266, 0.0122	1		
				6
Q4				
(i)	$x_{ij} = \mu + \alpha_i + e_{ij}$	1		
	$\mu$ = population	1		
	grand mean for whole experiment	1		
	$\alpha_i$ = population	1		
	mean by which <i>i</i> th treatment differs from			
	$\mu$ $e_{ii}$ are experimental errors	1		
		3	Allow "uncorrelated"	
	$\sim$ ind N $(0, \sigma^2)$		1 for ind N; 1 for 0; 1 for	
			$\sigma^2$ .	9
(ii)	Totals are 240, 246, 254, 264, 196			
	each from sample of size 5 Grand total 936			
	_			
	"Correction factor" CF = $\frac{936^2}{20}$ = 43804.8			
	20			
	Total SS = 44544 - CF = 739.2			
L		1	ı	

	Between contractors SS =				M1)	For correct methods for		
	$\frac{240^2}{5} + \dots + \frac{196^2}{5} - \text{CF} = 44209.6 - \text{CF} = 404.8$				M1J	any two,		
						A1	if each calculated SS is correct.	
	Residual SS	( by subt	traction) =	739.2 – 4	04.8 = 334.4	Λ'	correct.	
						-M1		
	Source of Variation	SS	df	MS	MS ratio	M1		
	Between					1		
	Contractors	404.8	3	134.93	6.456	A1		
	Residual	334.4	16	20.9				
					1		CAO	
	Total	739.2	19 🔍			1		
	Refer to F <sub>3,16</sub>				1	NO FT IF WRONG		
	Upper 5% point is 3.24				1	NO FT IF WRONG		
	Significant					1		
		ormance	es of con	tractors a	are not all	1		
	Seems performances of contractors are not all the same							
							12	
(iii)	Randomised blocks				B1			
	Description				E1	Take the subject areas as		
					E1	"blocks", ensure each contractor is used at least	3	
							once in each block	